

Japanese consumers' valuation of domestic beef after the Fukushima Daiichi Nuclear Power Plant accident

Abstract

After the radioactive contamination of agricultural and livestock products caused by the Fukushima Daiichi Nuclear Power Plant accident of March 11, 2011, consumer aversion against purchasing food products from the affected areas has become a major social problem in Japan. We examine how test results for radioactive materials in beef affect consumer valuation of beef produced in no-risk and affected areas using a choice experiment survey of consumers in the Tokyo metropolitan area ($N = 392$). Respondents were divided into two groups: one faced choice experiment tasks under the current test condition (the test status was only "under the limit"), and the other faced choice experiment tasks under the tightened test condition (with three levels: "below the limit," "below one-tenth of the limit," and "undetected"). We found that consumer valuation of "below the limit" beef in the affected area did not differ from that of "below one-tenth of the limit" beef in the affected area. Introducing the tightened status improved consumer valuations of all types of beef in the no-risk area regardless of the test status. However, consumer valuation of "undetected" beef in the affected area was lower than that in the no-risk area. The same measures need to be implemented with great care in both no-risk and affected areas, failing which the effects of measures taken in the affected areas may be diluted.

Keywords: radioactive contamination; domestic beef; ordered probit model; choice experiment; willingness-to-pay (WTP)

Introduction

The Great East Japan Earthquake and the subsequent tsunami on March 11, 2011 triggered the accident at the Tokyo Electric Power Company's Fukushima Daiichi Nuclear Power Plant, which released a massive amount of radioactive materials into the environment, contaminating the air, soil, and agricultural and livestock products (Baba, 2013). After limiting the maximum permissive radiation exposure from all food products to 5 millisievert (mSv) per year (for radioactive cesium), the Japanese government set a provisional limit for each food group on March 17, 2011. It intended to regulate the shipment and intake of specific agricultural and livestock products whenever radioactive material exceeding the limit was detected (see Note 1). Although the Food Safety Commission determined that food products meeting these limits had no health effects and declared them safe (Hamada & Ogino, 2012), reputational damage became a major issue; consumers and distributors who were concerned about

1 the risk of radioactive contamination of food refrained from purchasing agricultural and livestock
2 products from the affected area, as they suspected them to be contaminated. Further, some non-affected
3 production areas independently introduced tests to determine the presence of radioactive material in their
4 agricultural and livestock products and advertised their safety to consumers.

5 While the risk of nuclear effects is the greatest in Fukushima, it is almost non-existent for Kagoshima,
6 which is about 1,000 km away from Fukushima. We examine how the sociodemographic and
7 psychological characteristics of consumers affect their aversion to consuming beef from Fukushima and
8 Kagoshima Prefectures. Using a choice experiment, we test the effect of changes in labeling rules
9 depicting the status of radioactive testing on consumer valuation of beef by production area. Beef was
10 selected as the study subject; of all agricultural and livestock products, beef drew considerable social
11 interest because 1,530-2,700 becquerel (Bq) of radioactive cesium (exceeding the then limit of 500 Bq
12 per kilogram) was detected from the meat of 11 cows shipped from Fukushima in July 2011(MHLW,
13 2011). Moreover, the cause of this contamination, namely feeding cattle with rice straw containing highly
14 concentrated radioactive material, remains unprecedented for the general public.

15 Kuriyama (2012) and Ujiie (2011a; 2011b; 2012) studied consumer valuation of food contaminated
16 by this accident. Ujiie (2011a; 2011b; 2012) used the contingent valuation method to examine the
17 relationship between the results of radioactive material tests and consumer valuations of spinach, milk,
18 rice, and beef produced in Fukushima. Ujiie (2011a) estimated willingness to pay (WTP) for spinach
19 from Fukushima, assuming that spinach produced at a domestic place of production without fear of
20 contamination by a radioactive material is sold at JPY 150 per pack. Using a web survey conducted in
21 June 2011 of 392 married women living in the Tokyo metropolitan area, the mean WTP for spinach from
22 Fukushima containing below the limit, half of the limit or less, and no detected radioactive material, was
23 JPY 68, JPY 71, and JPY 98 per pack, respectively. Moreover, Ujiie (2011b) estimated the WTP for beef
24 from Fukushima, assuming that beef produced at a domestic place of production without fear of
25 contamination by a radioactive material is sold at JPY 200 per 100 g. Using a web survey of 868 married
26 women living in the same area in August 2011, the mean WTP for beef from Fukushima containing
27 below the limit, half of the limit or less, one-tenth of the limit or less, one-hundredth of the limit or less,
28 and no detected radioactive material was JPY 71, JPY 73, JPY 78, JPY 88, and JPY 118 per 100 g,
29 respectively. These studies demonstrated that consumers perceived agricultural and livestock products
30 labeled “contains radioactive material below the limit” and “contains radioactive material below half of
31 the limit” as being almost equivalent. These results also revealed that consumer valuations did not

1 improve unless the products were labeled “contains radioactive material below one-hundredth of the
2 limit” or “radioactive material was not detected.” Thus, reputational damage could be effectively
3 suppressed if detection levels were described in more detail and products were promoted directly to
4 consumers when radioactivity was undetected or was below one-hundredth of the limit. Kuriyama (2012)
5 used a choice experiment to examine consumer valuation of the level of radiation exposure from
6 consuming rice and found that consumers were willing to pay JPY 7 per kg for reducing the amount of
7 radiation exposure by 1 microsievert (μSv). He assumed that the relationship between the amount of
8 exposure and consumer valuation is linear. However, Ujiie (2011a; 2011b) demonstrated that it is highly
9 likely that consumers change their valuation according to the radiation dose.

10 These previous studies suffer from a limitation in that they did not consider a spillover effect, which is
11 the effect of information regarding a good/service on the evaluation of other goods/services not directly
12 referred to in the information (Ahluwalia, Unnava, & Burnkrant 2001; Hansen & Onozaka 2011; Roehm
13 & Tybout 2006). In our context, this refers to how the introduction of or changes in radioactive material
14 testing in agricultural and livestock products produced in non-affected areas—in an attempt to make
15 consumers feel safe—would affect their valuations of products produced in affected areas.

16 17 18 **Methods**

19 *Data*

20 A total of 412 respondents completed the web survey conducted between September 29 and October 1,
21 2011, which served as the source of the data for this study. The survey was administered by Macromill,
22 Inc., a leading Japanese online research company, which maintains a panel of more than 1.16 million
23 Japanese consumers. Our respondents met all of the following conditions: 1) resides in the Tokyo
24 metropolitan area (Tokyo, Chiba, Saitama, and Kanagawa Prefectures); 2) aged 20 years or older; 3)
25 purchased beef for grilling in the past six months; and 4) is the household member who purchases beef
26 most frequently. Panel members were invited to participate and the survey was closed once the desired
27 number of respondents was reached.

28 After excluding one respondent who left some questions unanswered and 19 respondents with
29 inconsistent cognition about the risks of consuming radioactively contaminated beef, the number of valid
30 respondents dropped to 392. As referred to hereinafter, there are two versions of the questionnaire, which
31 differ only in the scenario of choice experiment questions from the view point of labeling rules depicting
32 the status of the radioactive material test (see Stage 2: Choice experiments for details). There are 196

1 valid respondents in each version. The average age of the valid respondents was 43, which is younger
2 than the average age of 50 in the adult population in the area (SBJ, 2011). This is because the rate of
3 internet usage is lower among the older age group (MIAFC, 2011). Females accounted for 79% of all
4 valid respondents, which is significantly higher than the population average of 50%, based on the census
5 results. This is probably because the subjects who met the fourth sampling criterion are mostly women,
6 who are responsible for housework. The average annual household income among the valid respondents
7 was JPY 6.79 million, which is 14% higher than the corresponding value of JPY 5.96 million among
8 families living in the area in 2009 (SBJ, 2010).

9 The contents of the questionnaire were divided into six parts: anxiety for various food safety issues;
10 purchase intentions based on the status of radioactive material tests; choice experiments for valuing beef
11 according to the status of radioactive material tests; purchase experiences of meats before and after the
12 accident; knowledge of and attitudes toward food safety issues, including radioactively contaminated
13 food; and the respondent's individual and household characteristics.

14

15 ***Stage 1: Aversion to beef based on the status of radioactive material tests***

16 To investigate aversion to the beef based on the status of radioactive material tests, respondents were
17 asked how likely they were to purchase beef produced in Fukushima and Kagoshima when the
18 radioactive material in each beef product was “untested (not tested for radioactive materials),” “below
19 the limit (contains radioactive materials below the limit),” and “undetected (did not detect radioactive
20 materials).” They were asked to select an option that best described their opinion: “I would not purchase
21 beef even if it is cheap,” “I would purchase beef if it is comparatively cheap,” “I would purchase beef if
22 its price is reasonable,” and “I would purchase beef even if it is comparatively expensive” (see Appendix
23 A). Using combinations of responses to each beef category, 19 respondents were excluded from the
24 analysis because their cognition of the risk of contaminated beef was inconsistent. For example, a
25 response combination of “I would purchase ‘below the limit’ Fukushima beef if it is comparatively
26 cheap” and “I would not purchase ‘undetected’ Fukushima beef even if it is cheap” is inconsistent. A
27 similar comparison was performed for other beef categories to exclude inconsistent respondents. Then,
28 the responses were combined separately for Fukushima beef ($j = f$) and Kagoshima beef ($j = k$) into
29 ordinal variables with four categories (see Appendix A): 1) I would purchase beef produced in area j even
30 if it is “untested” ($Y_j = 0$); 2) I would not purchase “untested” beef produced in area j ($Y_j = 1$); 3) I would
31 not purchase beef produced in area j even if it shows “below the limit” radioactivity ($Y_j = 2$); and 4) I

1 would not purchase beef produced in area j even if radioactivity is “undetected” ($Y_j = 3$). Therefore, the
 2 categorical variable Y_j can be viewed as a measure of aversion to beef produced in area j based on the
 3 status of radioactive material tests. Y_j is used as the observable variable of the (latent) objective variable
 4 in the ordered probit analysis of aversion to beef and in the construction of the cutoff variables in the
 5 choice experiment analysis.

6
 7 **Stage 2: Choice experiments**

8 After answering questions about their purchase intentions, respondents faced the choice experiment
 9 questions, which presented them four options for beef produced in different areas. “None of these” was
 10 added as a fifth response option (Figure 1).

11
 12 Please select one of the following four types of beef that you would be most likely to purchase.
 13 If you would not purchase any of these types, please select “none of these.”
 14

| | Option 1 | Option 2 | Option 3 | Option 4 | Option 5 |
|-----------------------------------|-----------------------|-----------------------|-----------------------|-----------------------|-----------------------|
| Product origin | Fukushima beef | Kagoshima beef | Australian beef | U.S. beef | |
| Status of radioactive material | Below the limit | Below the limit | | | None of these |
| Price per 100 g | JPY 698 | JPY 798 | JPY 198 | JPY 198 | |
| Check only one circle | <input type="radio"/> | <input type="radio"/> | <input type="radio"/> | <input type="radio"/> | <input type="radio"/> |

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 16 **Figure 1** Sample choice experiment question (Scenario 1)

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Respondents were asked to select the beef they would be most likely to purchase. The options were differentiated along three attributes: product origin, radioactive material test result, and price (Table 1).

Table 1 Attributes and their levels.

| Attribute | Levels |
|--------------------------------|---|
| Product origin | Fukushima, Kagoshima, Australia, U.S. |
| Radioactive material testing | |
| Scenario 1 | Fukushima: below the limit Kagoshima: no label (untested), below the limit |
| Scenario 2 | Fukushima: below the limit, below one-tenth of the limit, undetected Kagoshima: no label (untested), undetected |
| Price per 100 g (Unit: JPY) | Fukushima: 298, 398, 498, 598, 698 Kagoshima: 398, 498, 598, 698, 798 Australian and U.S.: 98, 148, 198, 248, 298 |

Note: As of August 2013, USD 1 = JPY 97.5.

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Product origin was set as an alternative-specific attribute. Feeding cattle radioactively contaminated feed was a likely factor contributing to contamination of beef. This does not appear to have happened in Kagoshima, which is about 1,000 km away from Fukushima. Therefore, we selected the domestic production area of Kagoshima to compare against Fukushima. Given the realities of the beef market, the two domestic beef categories from both prefectures only include *Wagyu* breeds, and we set Australian and U.S. beef as the two imported beef categories. For constructing a choice situation similar to the actual purchase situation in a retail store, the respondents were informed that the only cut of beef referred to in the questions is the boned rib cut, which is suitable for *yakiniku*, a very popular type of Japanese cuisine involving grilling the meat with intense sauces. We provided no information regarding the quality of Australian and U.S. beef as these explanations are not usually provided at retail stores.

Two scenarios of the questionnaire, each with varying radioactive material test attribute settings, were created. Scenario 1 reflected the actual situation during the survey: all options for Fukushima beef were labeled “below the limit,” while those for Kagoshima beef were labeled either “below the limit” or “no label (untested)” (Figure 1). In response to the results of the tests conducted in July 2011, in which

1 radioactive cesium exceeding the limit was detected in Fukushima beef, all cattle shipped from
 2 Fukushima (and also cattle from several neighboring prefectures) are currently tested for radioactive
 3 material. Kagoshima and many other prefectures voluntarily test some of their cattle for radioactive
 4 material. Scenario 2, which includes additional labels of hypothetical test status, includes “below the
 5 limit,” “below one-tenth of the limit,” or “undetected” for Fukushima beef and “no label (untested)” or
 6 “undetected” for Kagoshima beef. Compared to Scenario 1, Scenario 2 provides additional details of the
 7 test results.

8 The price levels were determined based on the results of a market price survey conducted prior to this
 9 study and a previous study of choice experiments using beef as the subject (Aizaki, Sawada, Sato, &
 10 Kikkawa, 2012).

11 A *D*-efficiency-based approach (Zwerina, Huber, & Kuhfeld, 1996) was used for creating 10 choice
 12 experiment questions for each scenario, wherein 10 questions in scenario 1 differ from those in scenario
 13 2. All respondents faced 10 choice tasks in either scenario 1 or scenario 2. To reduce hypothetical bias,
 14 respondents were asked to read the cheap talk script before conducting these choice tasks (Van Loo,
 15 Caputo, Nayga, Metullenet, & Rick, 2011). The script is as follows: According to previous surveys,
 16 individuals’ willingness to pay for a good/service tends to be larger than the amount of money they
 17 actually pay for the same good/service in a store. This is because individuals tend to be lax about
 18 hypothetical spending decisions as they do not have to actually purchase the good/service. Please answer
 19 the following questions after reflecting the extent to which you may harbor such a tendency.

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 21 ***Statistical analysis***

22 *Ordered probit model analysis of purchase aversion*

23 The categorical variables Y_f and Y_k represent the respondent’s aversion to Fukushima beef and
 24 Kagoshima beef, respectively, and can be regarded as ordinals based on the respondent’s cognition of
 25 radiation risk. Therefore, the factors affecting aversion, according to the status of radioactive material
 26 tests, were examined using the following ordered probit model.

$$\begin{aligned}
 27 \quad Y_{ji}^* &= \alpha_{0j} + \alpha_{1j}FEM_i + \alpha_{2j}AGE_i + \alpha_{3j}CJS_i + \alpha_{4j}CES_i + \alpha_{5j}CPS_i \\
 28 \quad &+ \alpha_{6j}INC_i + \alpha_{7j}ARCF_i + \alpha_{8j}ASFD_i + \alpha_{9j}AAPE_i + \alpha_{10j}AAIE_i \\
 29 \quad &+ \alpha_{11j}FSI_i + \alpha_{12j}FSK_i + \alpha_{13j}FST_i + e_{ji} \quad j = f, k; i = 1, 2, \dots, N
 \end{aligned} \tag{1}$$

30 Respondent i ’s (unobservable) latent variable for production area j ($j = f, k$) is Y_{ji}^* . The latent variable
 31 and the observable variable Y_{ji} have the following relationship.

$$\begin{aligned}
1 \quad & Y_{ji} = 0 \quad \text{if} \quad Y_{ji}^* \leq 0 \\
2 \quad & = 1 \quad \text{if} \quad 0 < Y_{ji}^* \leq \mu_{1j} \\
3 \quad & = 2 \quad \text{if} \quad \mu_{1j} < Y_{ji}^* \leq \mu_{2j} \\
4 \quad & = 3 \quad \text{if} \quad \mu_{2j} < Y_{ji}^*
\end{aligned} \tag{2}$$

5 We further assume that the error term e_{ji} is normally distributed with mean 0 and variance 1. Parameter
6 vector α and threshold parameter vector μ can be estimated using the maximum-likelihood method
7 (Greene & Hensher, 2010).

8 Table 2 shows the definitions of the independent variables. It has been reported that sociodemographic
9 variables such as gender, age, the presence of children, and household income influence consumer
10 concerns about food safety and food purchasing behavior (Dosman, Adamowicz, & Hrudefy, 2001; Lin,
11 1995; Nayga, 1996). Moreover, in Japan, media reports and public announcements made by the
12 government have educated the public that the lifetime risk of dying of cancer caused by radiation
13 exposure is higher among *younger* children and that the risk among adults becomes lower as they age,
14 because their life expectancy becomes shorter (Gofman, 1990; Preston, Shimizu, Pierce, Suyama, &
15 Mabuchi, 2003; UNSCEAR, 1988).

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Table 2 Definition and descriptive statistics of independent variables used in the ordered probit analysis ($N = 392$).

| Variable | Definition | Mean | S.D. |
|-------------|---|------|-------|
| <i>FEM</i> | Dummy variable = 1 if the respondent is female, otherwise = 0 | 0.79 | 0.41 |
| <i>AGE</i> | Age of the respondent | 43.1 | 11.20 |
| <i>CJS</i> | Dummy variable = 1 if the respondent's family includes a junior high or/and high school student, otherwise = 0 | 0.16 | 0.37 |
| <i>CES</i> | Dummy variable = 1 if the respondent's family includes an elementary school student, otherwise = 0 | 0.15 | 0.36 |
| <i>CPS</i> | Dummy variable = 1 if the respondent's family includes a pre-elementary school student, otherwise = 0 | 0.17 | 0.38 |
| <i>INC</i> | Annual household income (Unit: million JPY) | 6.79 | 3.30 |
| <i>ARCF</i> | Degree of anxiety towards radioactively contaminated food ^a | 5.60 | 1.39 |
| <i>ASFD</i> | Degree of anxiety towards safety of feed ^a | 5.15 | 1.25 |
| <i>AAPE</i> | Degree of agreement with the statement: "I would like to support the affected area by purchasing food products produced there" ^b | 4.11 | 1.49 |
| <i>AAIE</i> | Degree of agreement with the statement: "Internal exposure by ingestion of food is inevitable to some extent now that radioactive materials have spread" ^b | 4.34 | 1.36 |
| <i>FSI</i> | Confidence in provided information about radioactive contamination of food ^c | 0.00 | 0.97 |
| <i>FSK</i> | Knowledge about radioactive material and countermeasures to contain its spread in the food supply ^c | 0.00 | 0.91 |
| <i>FST</i> | Confidence in the actions taken by the government and producers ^c | 0.00 | 0.93 |

Notes: ^a These items are measured on a 7-point scale, where 1 means "No anxiety," and 7, "Very high anxiety."

^b These items are measured on a 7-point scale, where 1 means "I completely disagree," and 7, "I completely agree."

^c These are normalized factor scores.

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5 For these reasons, we adopted *FEM*, *AGE*, *CJS*, *CES*, *CPS*, and *INC* as explanatory variables that may
6 influence the respondent's aversion to Fukushima beef and Kagoshima beef. In particular, three kinds of
7 "the presence of children"—preschool children (age 5 and below) dummy *CPS*, elementary school
8 student (6 to 12 years) dummy *CES*, and junior or/and senior high school student (13 to 18 years) dummy
9 *CJS*—were adopted in order to examine which age group of children influenced the aversion to beef
10 produced in these areas. Previous research has indicated that consumers' acceptance of and purchase
11 intentions for a particular food are also affected by their risk perception (anxiety) of the food, knowledge

1 and attitudes, trust in sources of information, and government action (Chen, 2007; Christoph, Bruhn, &
2 Roosen, 2008; McCarthy, O'Reilly, Cotter, & de Boer, 2004). Therefore, we adopt *ARCF*, *ASFD*, *AAPE*,
3 *AAIE*, *FSK*, *FSI*, and *FST* as explanatory variables reflecting the respondent's degree of anxiety towards
4 the radioactive contamination of food, the respondent's degree of anxiety towards the safety of the feed,
5 desire to help the affected area by purchasing food products produced there, respondent's resignation to
6 the fact that internal exposure to radioactive materials via ingestion of food is inevitable to some extent
7 now that the materials have spread, knowledge about radioactive material and countermeasures against
8 their spread in the food supply, confidence in the information conveyed to them about the radioactive
9 contamination of food, and confidence in the actions being taken by the government and producers to
10 address this contamination, respectively, which may affect the respondent's aversion to Fukushima beef
11 and Kagoshima beef. Here, *FSI*, *FSK*, and *FST* are the scores of three factors extracted from a factor
12 analysis of respondents' knowledge of and attitudes concerning radioactive contamination of food (see
13 Appendix B).

14 The Pearson's correlation coefficients for each pair of independent variables do not exceed the scope
15 of a reasonable limit (0.7) in terms of absolute value. The highest correlation coefficient (0.669) is
16 observed for *FSI* and *FST*. The second-highest correlation coefficient (0.642) is found for *FST* and *AAPE*.
17 All other correlation coefficients are below 0.5 in terms of absolute value. Therefore, we believe that a
18 serious multicollinearity problem will not arise.

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20 *Non-compensatory choice model analysis of choice experiment tasks*

21 In order to incorporate respondents' aversion to beef according to the status of radioactive material test
22 into the modeling of their decision-making in the choice experiment questions, we use a
23 non-compensatory choice model proposed by Swait (2001). This model is used to determine whether a
24 consumer applies non-compensatory or compensatory rules while valuing the attribute variables in the
25 utility. For example, in the context of the status of radioactive material tests, a respondents' decision "I
26 would not purchase beef if it is untested" can be viewed as one of the non-compensatory rules of
27 attributes. Thus, this model is suitable for our study.

28 Firstly, we construct a linear compensatory utility model. Respondent n is assumed to select one
29 alternative from a choice set containing four beef alternatives and an opt-out option. Under a linear
30 compensatory utility model, the systematic component of the utility function of option j , as Fukushima
31 beef ($j = f$), Kagoshima beef ($j = k$), Australian beef ($j = a$), U.S. beef ($j = u$), or none of these ($j = \text{nop}$), is

1 as follows.

2 Scenario 1:

$$\begin{aligned} 3 \quad V_f^c &= \beta_{0f} + \beta_{pf} P_f \\ 4 \quad V_k^c &= \beta_{0k} + \beta_{4k} RI_k + \beta_{pk} P_k \\ 5 \quad V_a^c &= \beta_{0a} + \beta_{pa} P_a \\ 6 \quad V_u^c &= \beta_{0u} + \beta_{pu} P_u \\ 7 \quad V_{nop}^c &= 0 \end{aligned} \tag{3}$$

8 Scenario 2:

$$\begin{aligned} 9 \quad V_f^c &= \gamma_{0f} + \gamma_{4f} R2_f + \gamma_{8f} R3_f + \gamma_{pf} P_f \\ 10 \quad V_k^c &= \gamma_{0k} + \gamma_{4k} R3_k + \gamma_{pk} P_k \\ 11 \quad V_a^c &= \gamma_{0a} + \gamma_{pa} P_a \\ 12 \quad V_u^c &= \gamma_{0u} + \gamma_{pu} P_u \\ 13 \quad V_{nop}^c &= 0 \end{aligned} \tag{4}$$

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15 where RI_k , $R2_f$, and $R3_j$ ($j = f, k$) are alternative-specific dummy variables for the label showing the
16 status of radioactive material tests: $RI_k = 1$, $R2_f = 1$, and $R3_j = 1$ indicate that the Kagoshima beef was
17 labeled “below the limit,” the Fukushima beef was labeled “below one-tenth of the limit,” and the beef
18 produced in area j was labeled “undetected.” Else, the variables take the value 0; P_j ($j = f, k, a, u$) denotes
19 alternative-specific attribute variables showing the price of beef produced in area j (unit: JPY per 100 g);
20 β_{0j} ($j = f, k, a, u$) are alternative specific constants; and the other β s and γ s are coefficients to be
21 estimated.

22 The non-compensatory choice model permits us to assume that while respondents have (un)acceptable
23 conditions of attributes (i.e., rules), such rules are occasionally violated. For example, although some
24 respondents may decide, as a rule, to purchase only domestic beef, they may occasionally violate it and
25 purchase imported beef. Under the non-compensatory choice model proposed by Swait (2001), these
26 rules and violations are expressed in the utility function using special dummy variables (i.e., cutoff
27 variables) and their estimated coefficients. Let us consider the rule that a respondent *does not* purchase
28 beef that has an undesirable feature for him/her (all the rules in our study are of this type), e.g., the rule in
29 this case would be “I do not purchase untested beef.” The cutoff variable corresponding to the rule takes
30 the value 1 if the respondent chooses to follow the rule and if the beef that is included in the respondent’s
31 choice set has a feature undesirable to him/her, and 0 otherwise. Thus, it should be noted that the cutoff

1 variable depends on both the beef attribute and the respondent's rule. Since the cutoff variable reflects
 2 the respondent's attitude toward an undesirable feature of the beef, the coefficient of the cutoff variable is
 3 expected to be negative. If the concerned respondent adheres strictly to the rule and the extent of
 4 undesirability is extremely large for him/her, the estimated coefficient of the cutoff variable would be an
 5 extremely large negative value. If the respondent does not adopt the rule while making his/her choice, the
 6 estimated coefficient would be a small negative value or 0.

7 We introduce two categories of cutoff variables: respondents' experiences of purchasing beef
 8 according to product origin and respondents' aversion to beef according to the status of radioactive
 9 material tests. Since some Japanese consumers are averse to specific product origins for a particular food
 10 product because of safety concerns or its taste (e.g., Aizaki, Sawada, Sato, & Kikkawa, 2012; Peterson &
 11 Yoshida, 2004), we create dummy variables NPE_w , NPE_a , and NPE_u , representing the *no* purchase
 12 experience of domestic beef, Australian beef, and U.S. beef, respectively. Here, 1 indicates that the
 13 respondent had *not* purchased beef of the respective product origin in the past, and 0, otherwise. We can
 14 consider a respondent's no-purchase experience of beef as his/her revealed rule of aversion to purchasing
 15 beef according to product origin. The second category of cutoff variables includes $DY1_j$, $DY2_j$, and $DY3_j$
 16 ($j = f, k$), which are dummy variables representing the aversion to purchasing beef according to the status
 17 of radioactive material tests. $DY1_j = 1$, $DY2_j = 1$, and $DY3_j = 1$ indicate that the respondent shows
 18 aversion to "untested" beef produced in area j , to "under the limit" beef produced in area j , and (even) to
 19 "undetected" beef produced in area j , respectively. Else, the variables take the value 0, which means that
 20 the respondent has no aversion to beef produced in area j . We can consider the respondent's intention not
 21 to purchase beef as his/her stated rule of aversion to purchasing beef according to the status of
 22 radioactive material tests.

23 Under the non-compensatory choice model and the cutoff variables mentioned above, the systematic
 24 component of the utility function of each option is given as:

25 Scenario 1:

$$\begin{aligned}
 26 \quad V_f^{nc} &= \beta_{0f} + \beta_{nw} NPE_w + \beta_{1f} DY1_f + \beta_{2f} DY2_f + \beta_{3f} DY3_f + \beta_{pf} P_f \\
 27 \quad V_k^{nc} &= \beta_{0k} + \beta_{nw} NPE_w + \beta_{1k} DY1_k + \beta_{2k} DY2_k + \beta_{3k} DY3_k + \beta_{4k} RI_k + \beta_{5k} RI_k \times DY1_k \\
 28 \quad &+ \beta_{6k} RI_k \times DY2_k + \beta_{7k} RI_k \times DY3_k + \beta_{pk} P_k \tag{5} \\
 29 \quad V_a^{nc} &= \beta_{0a} + \beta_{na} NPE_a + \beta_{pa} P_a \\
 30 \quad V_u^{nc} &= \beta_{0u} + \beta_{nu} NPE_u + \beta_{pu} P_u \\
 31 \quad V_{nop}^{nc} &= 0
 \end{aligned}$$

1 Scenario 2:

$$\begin{aligned}
2 \quad V_f^{nc} &= \gamma_{0f} + \gamma_{nw} NPE_w + \gamma_{1f} DY1_f + \gamma_{2f} DY2_f + \gamma_{3f} DY3_f + \gamma_{4f} R2_f + \gamma_{5f} R2_f \times DY1_f \\
3 \quad &+ \gamma_{6f} R2_f \times DY2_f + \gamma_{7f} R2_f \times DY3_f + \gamma_{8f} R3_f + \gamma_{9f} R3_f \times DY1_f \\
4 \quad &+ \gamma_{10f} R3_f \times DY2_f + \gamma_{11f} R3_f \times DY3_f + \gamma_{pf} P_f \\
5 \quad V_k^{nc} &= \gamma_{0k} + \gamma_{nw} NPE_w + \gamma_{1k} DY1_k + \gamma_{2k} DY2_k + \gamma_{3k} DY3_k + \gamma_{4k} R3_k + \gamma_{5k} R3_k \times DY1_k \\
6 \quad &+ \gamma_{6k} R3_k \times DY2_k + \gamma_{7k} R3_k \times DY3_k + \gamma_{pk} P_k \tag{6} \\
7 \quad V_a^{nc} &= \gamma_{0a} + \gamma_{na} NPE_a + \gamma_{pa} P_a \\
8 \quad V_u^{nc} &= \gamma_{0u} + \gamma_{nu} NPE_u + \gamma_{pu} P_u \\
9 \quad V_{nop}^{nc} &= 0
\end{aligned}$$

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11 Since the non-compensatory choice model can be integrated into any discrete choice models, we use the
12 error components multinomial logit (ECMNL) model to analyze the responses to the choice experiment
13 tasks. It is a type of flexible discrete choice model incorporating other probability terms into a
14 multinomial logit (MNL) model, to relax the assumption of the independence of irrelevant alternative
15 that exists within the MNL model (Brownstone, Bunch, & Train, 2000). This model has been applied in
16 recent empirical studies using choice experiments (MacDonald, Morrison, Rose, & Boyle, 2011;
17 Marcucci & Gatta, 2011). Under the ECMNL model integrated with the non-compensatory choice model,
18 the utility function of each option is specified as follows:

19 Scenario 1:

$$\begin{aligned}
20 \quad U_f^{nc} &= V_f^{nc} + \varepsilon_f + \theta_f E_f \\
21 \quad U_k^{nc} &= V_k^{nc} + \varepsilon_k + \theta_k E_k \\
22 \quad U_a^{nc} &= V_a^{nc} + \varepsilon_a + \theta_a E_a \tag{7} \\
23 \quad U_u^{nc} &= V_u^{nc} + \varepsilon_u + \theta_u E_u \\
24 \quad U_{nop}^{nc} &= \varepsilon_{nop} + \theta_{nop} E_{nop}
\end{aligned}$$

25 Scenario 2:

$$\begin{aligned}
26 \quad U_f^{nc} &= V_f^{nc} + \xi_f + \omega_f E_f \\
27 \quad U_k^{nc} &= V_k^{nc} + \xi_k + \omega_k E_k \\
28 \quad U_a^{nc} &= V_a^{nc} + \xi_a + \omega_a E_a \tag{8} \\
29 \quad U_u^{nc} &= V_u^{nc} + \xi_u + \omega_u E_u \\
30 \quad U_{nop}^{nc} &= \xi_{nop} + \omega_{nop} E_{nop}
\end{aligned}$$

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1 where ε_j is independent and identically type I extreme value distributed. E_j is an error component that
 2 assumes random respondent effects specific to choice j and averages to 0. The parameter vectors β , γ , θ ,
 3 and ω in each ECMNL model are estimated using the maximum simulated likelihood method (Train,
 4 2003). The estimations are conducted using NLOGIT 5.0 (Econometric Software, 2012).

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7 **Results**

8 *Ordered probit analysis*

9 Table 3 shows the respondents' aversion to Fukushima and Kagoshima beef according to the status of
 10 radioactive material tests. A high percentage (81%) showed aversion to "untested" Fukushima beef, that
 11 is, they stated that they would not purchase it. While this figure drops to 33% when Fukushima beef was
 12 labeled "below the limit," 25% showed aversion even "undetected" Fukushima beef, suggesting that
 13 many consumers have a strong fear of radioactively contaminated beef. Conversely, 23% showed
 14 aversion to "untested" Kagoshima beef. The figure declines significantly (to 6%) when Kagoshima beef
 15 was labeled "below the limit" and is 3% when it was labeled "undetected."

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Table 3 Respondents' aversion to Fukushima and Kagoshima beef according to the test status ($N = 392$).

| Test status | Averse to Fukushima beef | | Averse to Kagoshima beef | |
|-----------------|--------------------------|----|--------------------------|----|
| | <i>n</i> | % | <i>n</i> | % |
| Untested | 317 | 81 | 92 | 23 |
| Below the limit | 129 | 33 | 22 | 6 |
| Undetected | 99 | 25 | 11 | 3 |

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20 Tables 4 and 5 show the results of the ordered probit model analysis and the marginal effect of each
 21 independent variable calculated from the results, respectively.

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Table 4 Ordered probit model analysis of respondents' aversion to Fukushima beef according to the status of radioactive material tests ($N = 392$).

| Variable | Coefficient | S.E. | Marginal effects at means | | | |
|-------------------------|-------------|-------|---------------------------|-------------------|-------------------|-------------------|
| | | | Prob[$Y_i = 0$] | Prob[$Y_i = 1$] | Prob[$Y_i = 2$] | Prob[$Y_i = 3$] |
| Constant | 0.424 | 0.521 | | | | |
| <i>FEM</i> ^a | 0.271 * | 0.161 | -0.050 | -0.030 ** | 0.021 | 0.058 * |
| <i>AGE</i> | -0.016 *** | 0.006 | 0.003 ** | 0.002 ** | -0.001 ** | -0.004 ** |
| <i>CJS</i> ^a | 0.122 | 0.178 | -0.019 | -0.020 | 0.009 | 0.030 |
| <i>CES</i> ^a | 0.045 | 0.173 | -0.007 | -0.007 | 0.003 | 0.011 |
| <i>CPS</i> ^a | 0.334 * | 0.176 | -0.048 ** | -0.063 | 0.025 * | 0.087 * |
| <i>INC</i> | 0.031 | 0.019 | 0.007 | -0.005 | 0.002 | 0.007 |
| <i>ARCF</i> | 0.495 *** | 0.064 | -0.083 *** | -0.072 ** | 0.039 *** | 0.116 *** |
| <i>ASFD</i> | -0.083 | 0.065 | 0.014 | 0.012 | -0.007 | -0.020 |
| <i>AAPE</i> | -0.257 *** | 0.057 | 0.043 *** | 0.037 *** | -0.020 *** | -0.060 *** |
| <i>AAIE</i> | -0.046 | 0.049 | 0.008 | 0.007 | -0.004 | -0.011 |
| <i>FSI</i> | 0.207 ** | 0.090 | -0.034 ** | -0.029 ** | 0.016 ** | 0.047 ** |
| <i>FSK</i> | 0.120 * | 0.123 | -0.018 | -0.016 | 0.009 | 0.026 |
| <i>FST</i> | -0.484 *** | 0.110 | 0.079 *** | 0.069 *** | -0.037 *** | -0.111 *** |
| Threshold parameter | | | | | | |
| μ_1 | 2.020 *** | 0.079 | | | | |
| μ_2 | 2.354 *** | 0.083 | | | | |

Notes: $\chi^2_{13} = 254.97$ ***, McFadden's $R^2 = 0.240$.

For Tables 4 and 5, *** denotes $p < 0.01$, ** denotes $p < 0.05$, and * denotes $p < 0.10$.

^a For Tables 4 and 5, the marginal effects due to dummy variables are analyzed by taking the difference of estimated probabilities between the different levels of the dummy covariates.

Overall, females (*FEM*) were more likely to be averse to Fukushima beef than males. The probability of females refusing to purchase “undetected” Fukushima beef was higher. Females showed a similar tendency for Kagoshima beef. For example, the probability of them showing no aversion to “untested” Kagoshima beef was significantly lower than that of the males. Overall, aversion to Fukushima beef became significantly weaker with aging (*AGE*). The probability of showing no aversion to “untested” Fukushima beef was significantly higher, and the probability to refuse to purchase “undetected” Fukushima beef was significantly lower with aging. Respondents with preschool children (*CPS*) showed stronger aversion to Fukushima beef than others. The probability of them having no aversion to “untested” Fukushima beef was significantly lower, and the probability of refusing to purchase “undetected” Fukushima beef was higher.

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Table 5 Ordered probit model analysis of respondents' aversion to Kagoshima beef according to the status of radioactive material tests ($N = 392$).

| Variable | Coefficient | S.E. | Marginal effects at means | | | |
|-------------------------|-------------|-------|---------------------------|-------------------|-------------------|-------------------|
| | | | Prob[$Y_f = 0$] | Prob[$Y_f = 1$] | Prob[$Y_f = 2$] | Prob[$Y_f = 3$] |
| Constant | -1.997 *** | 0.662 | | | | |
| <i>FEM</i> ^a | 0.400 * | 0.205 | -0.102 ** | 0.077 ** | 0.013 ** | 0.012 * |
| <i>AGE</i> | -0.006 | 0.007 | 0.002 | -0.001 | 0.000 | 0.000 |
| <i>CJS</i> ^a | 0.196 | 0.203 | -0.059 | 0.042 | 0.009 | 0.008 |
| <i>CES</i> ^a | -0.223 | 0.208 | 0.059 | -0.044 | -0.008 | -0.007 |
| <i>CPS</i> ^a | 0.178 | 0.198 | -0.053 | 0.038 | 0.008 | 0.007 |
| <i>INC</i> | 0.011 | 0.022 | -0.003 | 0.002 | 0.000 | 0.000 |
| <i>ARCF</i> | 0.072 | 0.077 | -0.020 | 0.015 | 0.003 | 0.003 |
| <i>ASFD</i> | 0.225 *** | 0.079 | -0.064 *** | 0.047 *** | 0.009 ** | 0.008 ** |
| <i>AAPE</i> | -0.025 | 0.064 | 0.007 | -0.005 | -0.001 | -0.001 |
| <i>AAIE</i> | -0.106 ** | 0.054 | 0.030 ** | -0.022 ** | -0.004 * | -0.004 * |
| <i>FSI</i> | 0.067 | 0.102 | -0.019 | 0.014 | 0.003 | 0.002 |
| <i>FSK</i> | 0.192 ** | 0.085 | -0.055 ** | 0.040 ** | 0.008 ** | 0.007 ** |
| <i>FST</i> | -0.170 | 0.125 | 0.048 | -0.035 | -0.007 | -0.006 |
| Threshold parameter | | | | | | |
| μ_1 | 0.994 *** | 0.104 | | | | |
| μ_2 | 1.358 *** | 0.136 | | | | |

Notes: $\chi^2_{13} = 53.63$ ***, McFadden's $R^2 = 0.096$.

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Anxiety toward radioactively contaminated food (*ARCF*) significantly increased aversion to Fukushima beef. Larger *ARCF* values lowered the probability of showing no aversion to “untested” Fukushima beef and raised the probability of refusing “undetected” Fukushima beef. Conversely, respondents' aversion to Kagoshima beef was affected by their anxiety toward the safety of the feed (*ASFD*). The likelihood of showing no aversion to “untested” Kagoshima beef was lower among individuals with greater anxiety toward the safety of the feed. This is probably because the mass media would have alerted the public that radiation-contaminated feed was circulating across a wide area beyond the affected area. Further, individuals with a stronger desire to help the affected area by purchasing food products produced there (*AAPE*) showed weak aversion to “untested” Fukushima beef and had lower probability of refusing to purchase “undetected” Fukushima beef. Individuals who have resigned themselves to the fact that internal exposure by ingestion of food is inevitable due to the spread of radioactive materials in the food chain (*AAIE*) were relatively more likely to show no aversion to “untested” Kagoshima beef. *AAIE* had no observable effect on the aversion to Fukushima beef.

1 Individuals expressing higher confidence in the information provided by various sources on the
2 radioactive contamination of food (*FSI*) had stronger aversion to Fukushima beef. They were less likely
3 to show no aversion to “untested” Fukushima beef and more likely to refuse to purchase “undetected”
4 Fukushima beef. The information about the safety of food produced in Fukushima area presented through
5 various media at the time of the survey varied; some of it was positive (reassuring the respondents),
6 while some of it was negative (creating anxiety among them). Negative information was the primary
7 factor in decision making by the respondents; this is probably because people tend to err on the side of
8 caution. Individuals who have confidence in the actions taken by the government and producers to
9 address radioactive contamination in food (*FST*) showed significantly less aversion to Fukushima beef.
10 Given that knowledge on radioactive material and countermeasures against their spread to the food
11 supply (*FSK*) did not show a significant effect on the aversion, it appears that their aversion is more
12 likely to be influenced by the trustworthiness of the government’s/producers’ actions to address
13 radioactive contamination than their own knowledge. While *FSI* and *FST* showed no significant effect on
14 the aversion to Kagoshima beef, higher *FSK* confirmed significant aversion.

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16 ***Choice experiment analysis***

17 Table 6 shows the frequency of selection for each option in the choice experiment questions in
18 Scenarios 1 and 2. Of the five options, the relative frequency of selection was the highest for Australian
19 beef in both scenarios, and it was higher in Scenario 2 (46%) than in Scenario 1 (43%). Conversely, for
20 Fukushima beef (all types), the relative frequency of selection was lower in Scenario 2 (9%) than in
21 Scenario 1 (13%). This result suggests that describing additional details of radioactive material tests on
22 the label generally does not improve consumer valuation of Fukushima beef. In order to examine the
23 influence that each additional detail of the test result has on the consumer valuation of Fukushima beef,
24 let us look into the statistical analysis results on the response data from the choice experiments.

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Table 6 Frequency of selection for each option in the choice experiment tasks.

| Option | Scenario 1 | | Scenario 2 | |
|-----------------|------------|-----|------------|-----|
| | <i>n</i> | % | <i>n</i> | % |
| Fukushima beef | 245 | 13 | 174 | 9 |
| Kagoshima beef | 283 | 14 | 364 | 19 |
| Australian beef | 849 | 43 | 897 | 46 |
| U.S. beef | 365 | 19 | 382 | 19 |
| None of these | 218 | 11 | 143 | 7 |
| Total | 1,960 | 100 | 1,960 | 100 |

Table 7 shows the results of the ECMNL model analysis of the choice experiments. We exclude some interaction terms between attribute dummy variables (the status of radioactive material tests) and respondent characteristic variables (aversion to beef) from the final model because they produced unstable results. These unstable results may have occurred since the interaction terms that are not included in the final model rarely take the value 1.

Table 7 ECMNL model analysis of responses to choice experiment tasks.

| Variable | Scenario 1 | | Scenario 2 | |
|----------------------|-------------|-------|-------------|-------|
| | Coefficient | S.E. | Coefficient | S.E. |
| ASC_f | 9.679 *** | 0.845 | 7.105 *** | 0.870 |
| ASC_k | 5.609 *** | 0.763 | 8.420 *** | 0.749 |
| ASC_a | 6.158 *** | 0.450 | 5.979 *** | 0.463 |
| ASC_u | 5.750 *** | 0.506 | 5.956 *** | 0.492 |
| NPE_w | -5.354 *** | 0.981 | -4.195 *** | 0.658 |
| NPE_a | -6.715 *** | 0.738 | -4.964 *** | 0.719 |
| NPE_u | -6.571 *** | 1.457 | -5.091 *** | 0.594 |
| P_f | -1.331 *** | 0.093 | -1.174 *** | 0.126 |
| P_k | -1.089 *** | 0.077 | -1.269 *** | 0.095 |
| P_a | -0.864 *** | 0.073 | -1.029 *** | 0.091 |
| P_u | -1.559 *** | 0.117 | -1.434 *** | 0.120 |
| $DY1_f$ | -3.155 *** | 0.720 | -2.040 *** | 0.688 |
| $DY2_f$ | -5.547 *** | 1.387 | -6.558 *** | 2.121 |
| $DY3_f$ | -9.159 *** | 2.229 | -4.947 ** | 2.223 |
| $R2_f$ | - | | 0.302 | 0.339 |
| $R2_f \times DY3_f$ | - | | -1.761 | 1.859 |
| $R3_f$ | - | | 1.407 *** | 0.311 |
| $R3_f \times DY3_f$ | - | | -0.154 | 1.854 |
| $DY1_k$ | 1.137 | 0.947 | -1.191 | 0.773 |
| $DY2_k$ | 2.394 | 2.264 | -1.099 | 2.116 |
| $DY3_k$ | 2.764 ** | 1.382 | - | |
| RI_k | 1.126 *** | 0.190 | - | |
| $RI_k \times DY1_k$ | 0.671 * | 0.345 | - | |
| $RI_k \times DY2_k$ | -3.713 | 2.374 | - | |
| $RI_k \times DY3_k$ | -4.029 *** | 0.921 | - | |
| $R3_k$ | - | | 1.218 *** | 0.210 |
| Error component | | | | |
| E_f | 3.227 *** | 0.423 | 2.727 *** | 0.380 |
| E_k | 3.117 *** | 0.384 | 3.226 *** | 0.405 |
| E_a | 1.771 *** | 0.234 | 1.230 *** | 0.226 |
| E_u | 1.146 *** | 0.306 | 0.886 *** | 0.268 |
| E_{nop} | -3.225 *** | 0.431 | 2.402 *** | 0.403 |
| Log likelihood value | -1641.4 | | -1682.6 | |
| McFadden's R^2 | 0.480 | | 0.467 | |
| Replications | 500 | | 500 | |
| Sample size | 196 | | 196 | |
| Observations | 1960 | | 1960 | |

Notes: ASC_j is an alternative-specific constant for each option j (β_{0j} or γ_{0j}).

*** denotes $p < 0.01$, ** denotes $p < 0.05$, * denotes $p < 0.10$.

- : The applicable variable is not included in the model.

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The results in both questionnaire scenarios show that the estimated coefficient of the price of beef (P_j) was negative ($p < 0.01$): a price increase reduces the utility of beef. The dummy variable indicating whether a respondent has a purchasing habit (NPE_j) had a negative coefficient ($p < 0.01$); no experience of purchasing beef j in the past is associated with an enhanced likelihood of aversion to the applicable beef.

Consider the attributes of the radioactive material test. Since all the Fukushima beef was “below the limit” in Scenario 1, the results that each coefficient of $DY1_f$, $DY2_f$, and $DY3_f$ is significantly negative, meaning that respondents averse to “untested” or “below the limit” Fukushima beef or who show aversion to even “undetected” Fukushima beef, place less value on “below the limit” Fukushima beef compared to those who have no aversion to “untested” Fukushima beef. We applied “untested” and “below the limit” levels for Kagoshima beef in Scenario 1. Respondents averse to “untested” or “below the limit” Kagoshima beef have similar preferences to those without aversion to “untested” Kagoshima beef, because the coefficients of $DY1_k$ and $DY2_k$ are not significantly different from 0. The coefficient of $DY3_k$ is significantly positive; the choice probability of “untested” Kagoshima beef for those averse to even “undetected” Kagoshima beef is larger than that for others. Such increased choice probability seems to be inconsistent with their attitude toward the test status and could have resulted because only 11 respondents (3% of the total) show aversion to “undetected” Kagoshima beef. The main effect of RI_k was significantly positive (1.126, $p < 0.01$), implying that the “below the limit” label increases the utility level among respondents with zero interaction (respondents averse to Kagoshima beef even if it is “untested”). Of the respondents whose interaction was not 0, those averse to “untested” Kagoshima beef had a significantly positive ($p < .10$) interaction ($RI_k \times DY1_k$). However, those averse to even Kagoshima beef labeled “undetected” have a negative interaction ($RI_k \times DY3_k$), its absolute value (-4.029) being larger than the main effect (1.126). Thus, the additional value of “below the limit” is not recognized for those averse to even “undetected” Kagoshima beef.

The attribute of the radioactive material test in Scenario 2 included three levels for Fukushima beef—“below the limit,” “below one-tenth of the limit,” and “undetected”—and two levels for Kagoshima beef: “below the limit” and “undetected.” For Fukushima beef, whereas the main effect of “below one-tenth of the limit” ($R2_f$) is not significantly different from 0, that of “undetected” ($R3_f$) is significantly larger than 0. This means that consumer valuation of “below the limit” Fukushima beef is not different from that of “below one-tenth of the limit” Fukushima beef, and the valuation of

1 “undetected” Fukushima beef is larger than that of “below the limit” or “below the one-tenth of the limit”
2 Fukushima beef. For Kagoshima beef, the main effect of the “undetected” label ($R3_k$) was significantly
3 positive; the “undetected” label improves consumer utility more than the “untested” label.

4 Table 8 shows the representative respondent’s WTP for each beef type. “Representative respondents”
5 would be averse to buying “untested” Fukushima beef ($DYI_f = 1$), have no aversion to even “untested”
6 Kagoshima beef ($DYI_k = DY2_k = DY3_k = 0$), and have purchased U.S. beef, Australian beef, and Japanese
7 *Wagyu* beef ($NPE_w = NPE_a = NPE_u = 0$) in the past. WTP reveals how much representative respondents
8 are willing to pay in order to purchase the given beef rather than purchasing “none of these” (Lusk &
9 Schroeder, 2004). WTP per 100 g for “below the limit” Fukushima beef under Scenario 1 and Scenario 2
10 is JPY 490 and JPY 433, respectively. There is no statistical difference between these two values
11 according to the method of Poe, Giraud, & Loomis (2005). WTP for the “undetected” Fukushima beef
12 under Scenario 2 is JPY 553 per 100 g. Although the figure is significantly different from that of “below
13 the limit” Fukushima beef under Scenario 2, it is not significantly different from the WTP values for
14 “below one-tenth of the limit” Fukushima beef in Scenario 2 and “below the limit” Fukushima beef in
15 Scenario 1. Conversely, WTP per 100 g of “untested” and “below the limit” Kagoshima beef is JPY 515
16 and JPY 618, respectively, with no significant difference between the two. WTP per 100 g for “untested”
17 and “undetected” Kagoshima beef under Scenario 2 is JPY 664 and JPY 760, respectively, with a
18 statistically significant difference.

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2 **Table 8** Representative respondents' WTP for each beef type (Unit: JPY per 100 g).

| Beef type | Scenario | Product origin | Status of radioactive material | WTP | Significant pairwise differences in WTP by beef type |
|-----------|----------|-----------------------------|----------------------------------|----------------|---|
| (1) | 1 | Fukushima beef ^a | Below the limit | 490 [395, 587] | (1) > (5), (12) (1) < (3), (4), (9), (10),(11) |
| (2) | 1 | Kagoshima beef ^b | Untested | 515 [387, 646] | (2) > (5), (12) (2) < (4), (9), (10), (11) |
| (3) | 1 | Kagoshima beef ^b | Below the limit | 618 [497, 747] | (3) > (1), (5), (6), (7), (12) (3) < (10) |
| (4) | 1 | Australian beef | Untested | 645 [526, 777] | (4) > (1), (2), (5), (6), (7), (12) (4) < (10) |
| (5) | 1 | U.S. beef | Untested | 369 [312, 430] | (5) < (1), (2), (3), (4), (7), (8), (9), (10), (11) |
| (6) | 2 | Fukushima beef ^a | Below the limit | 433 [322, 551] | (6) < (3), (4), (8), (9), (10), (11) |
| (7) | 2 | Fukushima beef ^a | Below the one-tenth of the limit | 459 [351, 580] | (7) > (5) (7) < (3), (4), (9), (10), (11) |
| (8) | 2 | Fukushima beef ^a | Undetected | 553 [443, 675] | (8) > (5), (6), (12) (8) < (9), (10) |
| (9) | 2 | Kagoshima beef ^b | Untested | 664 [578, 754] | (9) > (1), (2), (5), (6),(7), (8), (12) (9) < (10) |
| (10) | 2 | Kagoshima beef ^b | Undetected | 760 [673, 852] | (10) > (1), (2), (3), (4), (5), (6), (7), (8), (9), (11), (12) |
| (11) | 2 | Australian beef | Untested | 583 [495, 684] | (11) > (1), (5), (6), (7), (12) (11) < (10) |
| (12) | 2 | U.S. beef | Untested | 416 [358, 481] | (12) < (1), (2), (3), (4), (8), (9), (10), (11) |

Notes: ^a Representative respondents' aversion to Fukushima beef is expressed as "I will purchase it if below the limit."

^b Representative respondents' aversion to Kagoshima beef is indicated as "I will purchase it even if untested."

Figures in brackets indicate the lower and upper values of the 95% confidence interval for each WTP, calculated according to the Krinsky & Robb (1986) procedure with 10,000 draws.

The Poe, Giraud, & Loomis (2005) one-sided combinatorial test was used to detect pairwise significant differences ($p < 0.10$) in WTP by beef type.

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4 WTP per 100 g of Australian beef under Scenario 1 and Scenario 2 is JPY 645 and JPY 583,
5 respectively, with no significant difference between the two. WTP for U.S. beef under Scenario 1 and
6 Scenario 2 is JPY 369 and JPY 416, respectively, with no significant difference between the two. WTP
7 for Australian beef is significantly larger than WTP for U.S. beef under both Scenario 1 and Scenario 2.
8 Detailing the status of radioactive material tests on domestic beef could not change the respondents'
9 valuation of imported beef. The WTP for Australian beef is significantly larger than "below the limit"

1 Fukushima beef and “untested” Kagoshima beef under Scenario 1. Also, it is significantly larger than
2 “below the limit” and “below the one-tenth of the limit” Fukushima beef under Scenario 2. Only WTP
3 for “undetected” Kagoshima beef under Scenario 2 is significantly larger than the WTP for Australian
4 beef. These findings suggest that the radioactive contamination accident may have decreased Japanese
5 consumers’ valuation of domestic *Wagyu* beef relative to Australian beef (Yoshida, 2013), which is
6 consistent with the result shown in Table 6.

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9 **Discussion and conclusion**

10 According to significant marginal effects, the ordered probit model analysis revealed that consumer
11 *anxiety towards radioactive contamination of food* and *confidence in the actions taken by the government*
12 *and producers* affected their aversion to Fukushima beef. Thus, to mitigate excessive risk aversion
13 behavior towards Fukushima beef, it is essential to win public trust by upscaling countermeasures taken
14 against radioactive contamination of food.

15 Twenty-three percent of respondents showed aversion to “untested” Kagoshima beef (Table 3).
16 Kagoshima Prefecture does not test its own agricultural and livestock products for radioactive material,
17 since there was no change in the amount of radiation in the environment before and after the accident.
18 Thus, it is highly unlikely that radioactive material could have contaminated the feed (see Note 2). The
19 fact that approximately one-fourth of respondents still showed a negative attitude towards Kagoshima
20 beef indicates the magnitude of the effect of this accident. However, the choice experiment did not
21 indicate that the attitude of not wanting to purchase “untested” Kagoshima beef had a significant impact
22 on deciding to purchase “untested” Kagoshima beef. Since the survey period coincided with the end of
23 September 2011, people may have had access to more information than they did immediately after
24 radioactively contaminated beef was discovered in early July 2011, thus making it easier for them to
25 evaluate the attributes of the beef. Even if the negative attitudes formed after the contamination was
26 discovered, we can conclude that the impact of discovering contaminated beef was small for beef
27 produced far from Fukushima (areas not directly affected by radioactivity).

28 Although about 80% of the respondents showed aversion to “untested” Fukushima beef, this figure
29 dropped to about 30% for “below the limit” Fukushima beef. This suggests that testing for radioactive
30 material plays a role in ensuring demand for Fukushima beef. However, the choice experiment results
31 suggest that describing additional details of radioactive material tests on the label generally does not help
32 reduce consumer anxiety about Fukushima beef (Table 6). Furthermore, the results of the choice

1 experiment analysis in Scenario 2 have elucidated the following points. The coefficient of “below
2 one-tenth of the limit” is not significantly different from 0. This implies that being “below one-tenth of
3 the limit” is considered the same as being “below the limit” from the viewpoint of easing consumer
4 anxiety. Although the limit of radioactive material per kilogram of beef was revised from 500 to 100
5 becquerel in October 2012, this result suggests that consumers probably experience very little benefit
6 from this revision. However, the “undetected” dummy variable had a significantly positive coefficient,
7 indicating that consumers deem it helpful in reducing anxiety. As Ujiie (2011a, 2011b) argued, we can
8 conclude that detailing the status of radioactive material tests on the label could possibly decrease
9 consumer anxiety about food safety. However, dividing the test status into three levels causes consumers
10 to differentiate among beef products, even though all three levels of beef produced in Fukushima meet
11 the safety standard for consumption. This implies that “below the limit” and “below one-tenth of the
12 limit” Fukushima beef could be rated relatively low by consumers, thus reducing its demand to the
13 detriment of cattle farmers. Therefore, if the government were to mandate food labels with more specific
14 test status, it is advisable to track changes in the market price of beef before and after such
15 implementation. It would be necessary to examine, as accurately as possible, whether such changes cause
16 losses to farmers and to compensate them accordingly.

17 We used Kagoshima beef as a competitor to Fukushima beef and differentiated Scenarios 1 and 2 by
18 changing the rules for the labeling status for testing radioactive materials (Table 1). This resulted in a
19 general relative increase in consumer valuation of Kagoshima beef: WTP per 100 g for “untested”
20 Kagoshima beef significantly increased in Scenario 2 (JPY 664) compared to Scenario 1 (JPY 515); WTP
21 per 100 g for “undetected” Kagoshima beef in Scenario 2 (JPY 760) was significantly higher than that for
22 “below the limit” Kagoshima beef in Scenario 1 (JPY 618). WTP per 100 g for “untested” Kagoshima
23 beef in Scenario 2 was significantly higher than those for all three types of Fukushima beef in Scenario 2
24 (JPY 433 to 553).

25 The following factors could have caused the abovementioned relationships. The majority of
26 respondents had no concerns about radioactive contamination of Kagoshima beef, and the tightened test
27 criteria by Kagoshima Prefecture reinforced this belief: beliefs about attributes of a product are one of the
28 constructs that determine consumer attitude toward the product (Fishbein & Ajzen, 1975). About 77% of
29 respondents showed no aversion to “untested” Kagoshima beef (Table 3); they trusted it would not have
30 been contaminated. This is because the difference between their WTP for “untested” and “below the
31 limit” Kagoshima beef was not significant (Table 8). For those without concerns about radioactive

1 contamination of Kagoshima beef, the “undetected” criterion introduced for the Kagoshima beef in
2 Scenario 2 also objectively proved their belief that “there is no risk of radioactive contamination to
3 Kagoshima beef.” Although the standard “undetected” applies only to the tested Kagoshima beef, it
4 could further reinforce consumer belief about the *overall* brand of Kagoshima beef; the rating of an
5 overall brand is influenced by the rating of its highest-rated product (Aker, 1991), “undetected”
6 Kagoshima beef. Therefore, it is conceivable that the tendency to choose Kagoshima beef increased
7 because the majority of the respondents improved their belief in the overall Kagoshima beef brand.

8 From what has been discussed above, we can conclude that the benefit of a measure taken for the
9 affected area may be diluted when a non-affected area employs the same measure. Thus, a situation could
10 be created wherein losses in the affected area cannot be quickly reduced/recovered as other production
11 areas adopt measures in their own interest. When a tremendous impact of the type seen here occurs in
12 limited regions, resolving the issue may take long if we rely on individual production areas to voluntarily
13 address it. Preferably, the central government should have adjusted the voluntary measures taken by each
14 production area for a specified period after the accident.

15 We treated the respondents’ aversion to purchase beef produced in the studied prefectures based on
16 the status of radioactive material tests as the dependent (endogenous) variable and exogenous variable in
17 the ordered probit analysis and the choice experimental analysis, respectively. Therefore, an endogeneity
18 bias in the estimates is possible (see Note 3). Because a small number of respondents were divided
19 between two separate questionnaire scenarios, it was impossible to apply flexible discrete choice models
20 to estimate the heterogeneity of preferences among respondents. However, our results are nevertheless
21 meaningful, because this is the first study to econometrically demonstrate Japanese consumers’ valuation
22 of domestic beef from both affected and non-affected areas based on test status for radioactive material.

23

24

25 **Notes**

26 (1) On April 1, 2012, the government further reduced the permissive dose for some foods to 1 mSv per
27 year. The revised limit for rice and beef came into effect on October 1, 2012.

28 (2) However, local shipping organizations and distributors test the beef produced in Kagoshima for
29 radioactive materials.

30 (3) One way to handle this issue is to follow the approach suggested by Ding, Veeman, & Adamowicz
31 (2012). We may estimate the ECMNL model using the predicted values for the dummy variables

1 related to the purchase intention according to the status of radioactive material tests included in
2 Equations (3) and (4). However, some dummy variables were predicted to be 0 for all respondents
3 (Greene & Hensher, 2010). Therefore, it was impossible to use the predicted values for estimating
4 our ECMNL model.

5
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32

1 **Appendix A.** Questions to gauge consumers' aversions to purchasing Fukushima and Kagoshima beef and
 2 definitions of the ordered categorical variables Y_f and Y_k

3 Listed below are six categories of beef for *yakiniku* (grilling). Please select the option most appropriate for you for
 4 each category. All the categories are *Wagyu* breeds.

| Categories of beef | Options | | | |
|--|--|---|---|---|
| | 1 I would purchase beef even if it is comparatively expensive | 2 I would purchase beef if its price is reasonable | 3 I would purchase beef if it is comparatively cheap | 4 I would not purchase beef even if it is cheap |
| A. Fukushima beef not tested for radioactive materials | <input type="radio"/> | <input type="radio"/> | <input type="radio"/> | <input type="radio"/> |
| B. Fukushima beef with radioactive materials below the limit | <input type="radio"/> | <input type="radio"/> | <input type="radio"/> | <input type="radio"/> |
| C. Fukushima beef without detectable radioactive materials | <input type="radio"/> | <input type="radio"/> | <input type="radio"/> | <input type="radio"/> |
| D. Kagoshima beef not tested for radioactive materials | <input type="radio"/> | <input type="radio"/> | <input type="radio"/> | <input type="radio"/> |
| E. Kagoshima beef with radioactive materials below the limit | <input type="radio"/> | <input type="radio"/> | <input type="radio"/> | <input type="radio"/> |
| F. Kagoshima beef without detectable radioactive materials | <input type="radio"/> | <input type="radio"/> | <input type="radio"/> | <input type="radio"/> |

5

6 Definition of Y_f

7 If respondent i selects

- 8 Any option other than 4 for categories A–C, then $Y_{fi} = 0$.
- 9 Option 4 for category A and any option other than 4 for categories B and C, then $Y_{fi} = 1$.
- 10 Option 4 for categories A and B and any option other than 4 for category C, then $Y_{fi} = 2$.
- 11 Option 4 for all categories A–C, then $Y_{fi} = 3$.

12 Definition of Y_k

13 If respondent i selects

- 14 Any option other than 4 for categories D–F, then $Y_{ki} = 0$.
- 15 Option 4 for category D and any option other than 4 for categories E and F, then $Y_{ki} = 1$.
- 16 Option 4 for categories D and E and any option other than 4 for category F, then $Y_{ki} = 2$.
- 17 Option 4 for all categories D–F, then $Y_{ki} = 3$.

18

Appendix B. Result of factor analysis assessing respondents' knowledge of and attitudes on radioactive contamination of food

A factor analysis was conducted to extract the respondents' knowledge of and attitudes on radioactive contamination of food from the responses to questions about 13 items concerning this problem. Respondents rated these items on a 7-point scale, where 1 means "I completely disagree/I have never heard of it" and 7 means "I completely agree/I have certainly heard of it." SPSS 12.0 (SPSS Inc., Chicago, IL, USA) with a principal factor method with promax rotation was used for the analysis. The interpretation of each of the extracted factors was based on factor loadings ≥ 0.50 .

The factor analysis of these statements extracted 3 factors as shown in the following table. The first of the three factors is highly correlated with reliability about the information related to radioactive contamination of food items provided by mass media, experts, and the government. This factor is labeled as *confidence in the information about radioactive contamination of food (FSI)*. The second factor, labeled as *knowledge on radioactive material and countermeasures against its spread in the food supply (FSK)*, is strongly associated with the knowledge about radioactive material and countermeasures against its spread in the food chain. Finally, the third factor is highly correlated with consumer confidence in the limit of radioactive materials in food- and shipment-related regulations mandated by the government, and it is positively associated with consumer trust in the efforts of the farmers to produce safe agricultural and livestock products. This factor is labeled *confidence in the actions taken by the government and producers (FST)*.

Factor loading matrix after promax rotation ($N = 392$).

| | Factor 1 | Factor 2 | Factor 3 |
|--|-------------|-------------|-------------|
| TV reports* are reliable | 0.96 | -0.01 | -0.05 |
| Newspaper reports* are reliable | 0.96 | -0.01 | -0.05 |
| Information* provided by experts is reliable | 0.69 | -0.02 | 0.08 |
| Information* published by the government is reliable | 0.58 | 0.05 | 0.28 |
| Following the Fukushima accident, the government promptly established limits for radioactive materials in food items | 0.03 | 0.70 | -0.24 |
| Even if cattle have ingested radioactive materials, providing them with clean feed will gradually decrease these materials inside their bodies | 0.02 | 0.69 | 0.12 |
| Occasional consumption of food items that exceed the limit poses no health risks | -0.03 | 0.66 | 0.11 |
| Since radioiodine has a short half-life, limits have not been established for radioiodine in meat for consumption | -0.08 | 0.63 | 0.19 |
| Different types of radiation have different effects on the human body | 0.03 | 0.61 | -0.20 |
| Some local governments are voluntarily testing food items other than those they are mandated to test for radioactive contamination | 0.01 | 0.58 | 0.01 |
| Tested food items containing radioactive materials below the limit are safe for consumption | -0.04 | -0.03 | 0.84 |
| Food items available on the market after being tested for radioactive materials and subject to government countermeasures are safe | 0.09 | 0.00 | 0.81 |
| Farmers take adequate care to provide safe agricultural and livestock products | 0.13 | 0.05 | 0.50 |

Notes: * implies that all reports/information refer to radioactive contamination of food items. Factor loadings are from factor pattern coefficients. Factor loadings ≥ 0.50 appear in bold.

